GENERICITY AND UD-RANDOM REALS

WESLEY CALVERT AND JOHANNA N.Y. FRANKLIN

Abstract. Avigad introduced the notion of UD-randomness based on Weyl’s 1916 definition of uniform distribution modulo one. We prove that there exists a weakly 1-random real that is neither UD-random nor weakly 1-generic. We also show that no 2-generic real can Turing compute a UD-random real.

1. Introduction

In [1] Avigad introduced a randomness notion based on uniform distribution modulo one, as described by Weyl in [10]. A sequence \( \langle x_i \rangle \) of real numbers is said to be uniformly distributed modulo one in the case that for every interval \( I \subseteq [0, 1] \), we have
\[
\lim_{n \to \infty} \frac{\left| \left\{ i < n \mid [x_i] \in I \right\} \right|}{n} = \mu(I)
\]
where \( \mu \) is Lebesgue measure and \([x_i]\) denotes the fractional part of \( x_i \). Weyl proved that if \( \langle a_i \rangle \) is a sequence of distinct integers, then for almost every real \( x \), the sequence \( \langle a_i x \rangle \) is uniformly distributed modulo one [10].

The definition below considers recursive sequences \( \langle a_i \rangle \) of distinct integers, observing that there are countably many such sequences and that therefore almost every \( x \in [0, 1] \) has the property that \( \langle a_i \rangle \) is uniformly distributed modulo one:

**Definition 1.1.** [1] A real \( x \in [0, 1] \) is UD-random if and only if for every recursive sequence \( \langle a_i \rangle \) of distinct integers, \( \langle a_i x \rangle \) is uniformly distributed modulo one.

We know several facts about the relationship between UD-randomness and other randomness notions, all due to Avigad’s original paper. Every Schnorr random real is UD-random; however, there are UD-random reals that are not weakly 1-random [1]. This is surprising, since weak 1-randomness is such a weak randomness notion: not only is every Schnorr random weakly 1-random, but so is every weakly 1-generic [7]. Avigad observed that, in consequence, there are weakly 1-random reals that are not UD-random, since a UD-random must satisfy the law of large numbers, and no real that is even weakly 1-generic does.

In this paper we investigate the relationship of the UD-random reals to the weakly 1-random reals and the weakly 1-generic reals further, and determine when a generic real can Turing compute a UD-random. There are very few conclusions that can be drawn about Turing degrees of UD-randoms from existing results. We can conclude, for instance, that every high Turing degree contains a UD-random because every high Turing degree contains a Schnorr random real [8] and that there is therefore a low UD-random and a hyperimmune-free UD-random using various basis
theorems. However, we cannot use any of our standard results involving the coincidence of randomness notions on given classes of Turing degrees. For instance, in the hyperimmune-free degrees, weak 1-randomness is equivalent to Martin-Löf-randomness [8], but we cannot conclude that this also holds of UD-randomness, since there are UD-randoms that are not weakly 1-random. We also know that Schnorr and Martin-Löf randomness coincide outside the high degrees [8], but we cannot conclude that UD-randomness coincides with these notions there as well.

We describe the weakly 1-random degrees further using UD-randomness in Section 2, and in Section 3 we discuss the relationship between the Turing degrees of UD-randoms and generic reals. Section 4 is dedicated to open questions.

1.1. Basic observations and notation. Our notation is standard and follows Soare [9] and Downey and Hirschfeldt [3]; we add that we denote the concatenation of two finite binary strings \( \sigma \) and \( \tau \) by \( \sigma \mathord{\cdot} \tau \).

Since we often discuss finite binary strings in the context of uniform distribution modulo one, we make the following convention: \( [a(\sigma)] \) is only defined if \( a < 2^{\vert \sigma \vert} \); that is, if \( \sigma \) is not so short that it is impossible for \( a(\sigma) \) to have a nonzero fractional part.

Fact 1.2. A real \( X \) is UD-random if and only if it is UD-random with respect to dyadic intervals.

Proof. One implication is obvious. For the other, suppose \( X \) is not UD-random. Suppose, in particular, that for an interval \( I \) and for arbitrarily large \( n \) we have
\[
\left| \frac{\vert \{ i < n \mid [x_i] \in I \} \vert}{n} - \mu(I) \right| > 2\epsilon.
\]
Now we can find \( I' \) such that \( I' \) has dyadic endpoints and \( \mu(I \triangle I') < \epsilon \). Then for arbitrarily large \( n \) we have
\[
\left| \frac{\vert \{ i < n \mid [x_i] \in I' \} \vert}{n} - \mu(I') \right| > \epsilon.
\]

\( \square \)

2. Categorizing weakly 1-random reals

Since the UD-random reals and the weakly 1-generic reals are disjoint, one can reasonably ask whether the UD-random weakly 1-random reals and the weakly 1-generic reals partition the weakly 1-random reals or whether there is a weakly 1-random real that is neither UD-random nor weakly 1-generic. We answer this question with the following theorem.

Theorem 2.1. There is a weakly 1-random real that is neither UD-random nor weakly 1-generic.

Proof. To construct such an \( X \in 2^\omega \), we must build not only \( X \) itself but also a dense \( \Sigma^0_1 \) set of strings \( S \) that \( X \) avoids (witnessing that \( X \) is not weakly 1-generic) and an infinite recursive sequence \( \langle a_i \rangle \) witnessing that \( X \) is not UD-random with respect to the interval \( (\frac{1}{2}, 1) \). We must also satisfy the following requirement for each \( \Sigma^0_1 \) set \( W_e \) to ensure that \( X \) is weakly 1-random:

\( R_e \): If \( \mu([W_e]) = 1 \), then \( X \in W_e \).

Without loss of generality, we take \( W_e \) to be a subset of \( 2^{<\omega} \) rather than \( \omega \).

This finite injury construction requires us to keep track of two sequences of finite binary strings for each \( e \): \( \langle \sigma_{e,s} \rangle_s \) and \( \langle \tau_{e,s} \rangle_s \). Each \( \tau_{e,s} \) is a string that is our guess for \( X \) under the assumption
that no requirement that has higher priority than $R_e$ will ever act after stage $s$, and the $\sigma_{e,s}$ are extensions of the $\tau_{e,s}$ that we construct in tandem with the $a_i$-sequence to ensure that $X$ will not be UD-random. We also build an auxiliary sequence $(n_{e,s})_s$ that will keep track of the measure we need to activate requirement $R_e$ at stage $s$, and we will also record the length of the $a_i$-sequence at the end of each stage and denote this by $\ell_s$.

In general, our strategy is this. At stage $s$, we have a list of the $e < s$ such that $R_e$ has already acted. For every other $e < s$, we will keep track of those $e' < e$ such that $R_{e'}$ has already acted, and we will have an approximation $\sigma_{e,s-1}$ to $X$ based on this list. If our list already includes all $e' < e$ such that $R_{e'}$ will act at any point, then our guess will be correct. We will work in the cylinder above $\sigma_{e,s-1}$ and wait for a lower-priority $W_e$ to cover a certain fraction of this cylinder. If this happens, we will act to satisfy $R_e$ and extend our $\sigma_{e,s-1}$ to a string $\tau_{e,s}$ to meet $W_e$. Then we will extend $\tau_{e,s}$ to $\sigma_{e,s}$ and extend our sequence of $a_i$s accordingly to make sure that $\frac{1}{2} < |a_i(\sigma_{e,s})| < 1$ for no more than one-fourth of all $a_i$ in our sequence. To build a dense set of strings $S$ that $X$ avoids, we will add strings to $S$ every time a new requirement $R_e$ acts to ensure that each string at a new, higher level is extended by an element of $S$. To this end, we will also keep track of the number of times we have acted to satisfy a requirement: we define $t_s$ to be this number at the end of each stage $s$.

**Construction.** Set $S_0 = \emptyset$ and $(a_i) = (\emptyset)$, and let $\sigma_{e,0} = (\emptyset)$ for all $e$, and $\ell_0 = t_0 = 0$.

At stage $s$, suppose we have a list $0 \leq e_0 < e_1 < e_2 < \ldots < e_k < s$ such that $R_{e_i}$ has acted already for each $i$. First, determine whether there is an $e$ less than $s$ such that

\begin{equation}
\mu([W_{e,s}] \cap [\sigma_{e,s-1}]) > \frac{1}{2} \mu([\sigma_{e,s-1}])
\end{equation}

and $R_e$ is not currently satisfied. If there is not, let $S_s = S_{s-1}$ and set $\sigma_{e,s} = \sigma_{e,s-1}$ for all $e$ and go on to the next stage. If there is such an $e$, choose the least one and act to satisfy $R_e$.

We begin by choosing a $\tau_{e,s}$ in $W_e$ that extends $\sigma_{e,s-1}$ and avoids $S_{s-1}$; we choose the leftmost shortest such string. If $a_i(\tau_{e,s})$ is not defined for some $i < \ell_{s-1}$, find the least $k$ such that $a_i < 2^k$ for each $i < \ell_{s-1}$ and extend $\tau_{e,s}$ by 0s until the extension has length $k$. For convenience, we refer to this extended string as $\tau_{e,s}$ from this point onwards.

Now $a_i(\tau_{e,s})$ must be defined for each $i < \ell_{e,s-1}$, so we check to see whether $\frac{|\{i < \ell_{s-1} \mid \frac{1}{2} < |a_i(\tau_{e,s})| < 1\}|}{\ell_{s-1}}$ is less than $\frac{1}{4}$. If this is the case, set $\sigma_{e,s} = \tau_{e,s}0$ and let $a_{\ell_{s-1}+1} = 2^m$, where $m$ is the position of this last 0. Then the fractional part of $a_{\ell_{s-1}+1}(\sigma_{e,s})$ is less than $\frac{1}{2}$, and, since $\ell_s = \ell_{s-1} + 1$, it follows that $\frac{|\{i < \ell_s \mid \frac{1}{2} < |a_i(\sigma_{e,s})| < 1\}|}{\ell_s} < \frac{1}{4}$ as well.

On the other hand, if $\frac{|\{i < \ell_{s-1} \mid \frac{1}{2} < |a_i(\tau_{e,s})| < 1\}|}{\ell_{s-1} + N} < \frac{1}{4}$, find the least $N$ such that $\frac{|\{i < \ell_{s-1} + N \mid \frac{1}{2} < |a_i(\tau_{e,s})| < 1\}|}{\ell_{s-1} + N} < \frac{1}{4}$. Take $\sigma_{e,s} = \tau_{e,s}0N$ and set $a_{\ell_{s-1}+i} = 2^{\ell_{s}+i}$ for all $1 \leq i \leq N$. For each $i$ such that $\ell_{s-1} < i \leq \ell_{s-1} + N$, the fractional part of $a_i(\sigma_{e,s})$ is less than $\frac{1}{4}$, so $\frac{|\{i < \ell_s \mid \frac{1}{2} < |a_i(\sigma_{e,s})| < 1\}|}{\ell_s} < \frac{1}{4}$ once more.

At this point we say that $R_e$ has been met. However, we are not done with stage $s$: we must extend our set $S$ to ensure its density. To do this, consider each string of length $|\sigma_{e,s}|$. If such a string has an extension in $S_{s-1}$, we need not extend it again; otherwise, extend it to a string long enough that the measure of $[\tau]$ is no more than $\frac{1}{2^{\ell_{s-1}+2}} \cdot \mu([\sigma_{e,s}])$. This guarantees that the
Lemma 2.3. Each requirement \( X = \lim_{e} \lim_{s} \sigma_{e,s} \) is no more than half the available measure.

At this point, we are ready to end this stage. We have already determined \( \sigma_{e,s} \). For \( e' > e \), reinitialize \( \sigma_{e',s} \) by setting it equal to \( \sigma_{e,s} \). If we acted to satisfy a requirement at this stage, define \( t_s = t_{s-1} + 1 \).

Let \( X = \lim_{e,s} \sigma_{e,s} \).

Verification.

Lemma 2.2. The limit \( X = \lim_{e} \lim_{s} \sigma_{e,s} \) exists.

Proof. We proceed by induction. Consider \( e \), and let \( s_0 \) be the first stage such that all \( e' < e \) that will ever act already have. If we never act to satisfy \( R_e \), then \( \sigma_{e,s_0} = \sigma_{e,t} \) for all \( t > s_0 \). If we act to satisfy \( R_e \) at a later stage \( s_1 \), then we never reinitialize \( \sigma_{e,s_1} \) and we have \( \sigma_{e,s_1} = \sigma_{e,t} \) for all \( t > s_1 \). Thus, \( \sigma_e = \lim_{e,s} \sigma_{e,s} \) exists. Since in our construction, whenever \( e_1 > e \), we have \( \sigma_e \prec \sigma_{e_1} \), the limit \( X = \lim_{e} \sigma_{e} \) exists as well.

Lemma 2.3. Each requirement \( R_e \) is satisfied and thus \( X \) is weakly 1-random.

Proof. If \( \mu([W_e]) \neq 1 \), then \( R_e \) is satisfied trivially. Suppose \( \mu([W_e]) = 1 \), and suppose that \( s_0 \) is the first stage such that all \( e' < e \) that will ever act already have. After \( s_0 \), \( \sigma_{e,s} \) will never be initialized again, so \( \sigma_{e,s_0} = \sigma_{e,t} \) for all \( t > s_0 \). Since \( \mu([W_e]) = 1 \), there will be some stage \( s_1 > s_0 \) such that \( \mu([W_{e,s}] \cap [\sigma_{e,s-1}]) > \frac{1}{2} \mu([\sigma_{e,s-1}]) \). At this stage, we act to satisfy \( R_e \) and will define \( \sigma_{e,s_1} \) so \( [\sigma_{e,s_1}] \) is contained in \( [W_e] \). After this stage, \( \sigma_{e,s} \) will never be initialized again, so we will have \( \sigma_e = \sigma_{e,s_1} \), and since \( \sigma_0 \prec \sigma_1 \prec \sigma_2 \prec \ldots \), we have \( X \in [W_e] \).

Lemma 2.4. \( X \) is not weakly 1-generic.

Proof. We first show that the set of strings \( S \) we constructed is a \( \Sigma^0_1 \) dense set as required. It is clear from the construction that it is \( \Sigma^0_1 \). Furthermore, at each stage \( s \) at which we add strings to \( S \), we extend all the strings of length \( |\sigma_{e,s}| \) for the appropriate \( e \). Since the lengths of the \( \sigma_{e,s} \) that we work with are unbounded, we will extend every element of \( 2^{<\omega} \) by an element in \( S \) at some point.

Now we show that \( X \) does not meet the \( \Sigma^0_1 \) set of strings \( S \) we constructed. We note that the measure of the strings above any \( \sigma_{e,s-1} \) is no more than (in fact, is strictly less than)

\[
\left( \frac{1}{2^2} + \frac{1}{2^3} + \ldots \right) \cdot \mu([\sigma_{e,s-1}]) = \frac{1}{2} \cdot \mu([\sigma_{e,s-1}]).
\]

Since at each stage at which we set new values of a \( \sigma_{e,s} \), we know that Inequality 1 holds, we can see that there is an string \( \tau \) extending an element of \( W_{e,s} \) such that \( [\tau] \cap [S_{s-1}] = \emptyset \), and we choose such a string as our \( \tau_{e,s} \). Therefore we have \( [\sigma_{e,s}] \cap [S_s] = \emptyset \) as well, and since every \( \sigma_e \) avoids \( S \), it follows that \( X \) avoids \( S \) as well.

Lemma 2.5. \( X \) is not UD-random.

Proof. We show that the sequence \( \langle a_i \rangle \) we built witnesses that \( X \) is not UD-random. First, we note that \( \langle a_i \rangle \) is recursive and infinite by construction: at every point at which we act to satisfy a requirement, we add at least one more element to this sequence, and since there are infinitely many \( W_e \) with measure 1, we act to satisfy a requirement infinitely often. It is clear that \( \langle a_i \rangle \) is recursive.
Lemma 2.2 shows us that \( X = \lim_{\sigma} e_{\sigma} \). Each time we act to meet a requirement \( R \), we also define an initial part of \( \langle a_i \rangle \) such that \([a_i(\sigma_{e,s})] = [a_i(\sigma_e)] = [a_i(X)]\) for all \( i < \ell \). Since the values of \( \ell \) associated with true values of \( \sigma_e \) form an unbounded sequence, we know that there are infinitely many \( \ell \) such that \( \left| \{ i < \ell \mid \frac{1}{2} < [a_i(X)] < 1 \} \right| < \frac{1}{4} \) and thus \( X \) is not UD-random with respect to the sequence \( \langle a_i \rangle \) and the interval \((\frac{1}{2}, 1)\).

This completes the proof of the Theorem.

Porism 2.6. If a real \( X \) is not UD-random, a recursive sequence can be found witnessing this for any interval \( I \subseteq [0, 1] \).

3. Genericity and UD-random reals

There has been a great deal of work relating, on the one hand, Turing degrees of Martin-Löf random, Schnorr random, and recursively random reals, and, on the other, the Turing degrees of generic reals [2, 5, 8].

Consider the following question: Given \( n \), is it possible for a (weakly) \( n \)-generic real to Turing compute a UD-random even though no (weakly) \( n \)-generic real can be UD-random itself? The answer is certainly yes for weak 1-generics, since every hyperimmune degree contains a weakly 1-generic real [7, 6], and there are Schnorr random reals that are hyperimmune. In fact, even a 1-generic real can be Turing equivalent to a UD-random real, since there is a high 1-generic.

Here, we provide a proof that this bound is tight: No 2-generic can Turing compute a UD-random real. Franklin proved in [4] that non-high 1-generics cannot Turing compute Schnorr random reals, which implies that no 2-generic can Turing compute a Schnorr random real. She also proved that no 1-generic can tt-compute a Schnorr random real. However, we cannot adapt the first argument to the UD-random case because it relies on the fact that outside the high degrees, Martin-Löf randomness and Schnorr randomness are equivalent. No such result is known for UD-randomness.

Theorem 3.1. No 2-generic real can Turing compute a UD-random real.

Proof. Let \( G \) be a 2-generic real, and let \( X \) be a non-recursive real that is Turing reducible to \( G \) via a Turing functional \( \Psi \). We wish to show that \( X \) is not UD-random. If \( X \) is recursive, then there is nothing to be done.

We begin by considering the following statements:
• \( \Psi^Z \) is total.
• There are infinitely many \( n \) for which \( \Psi^Z(n) = 0 \).

Each of these is a \( \Pi^0_2 \) statement, so \( G \) must force each of these statements to either be true or false. Since \( X = \Psi^G \), it follows that \( G \) must force the first to be true. Also, \( X \) is not recursive and thus not cofinite, so that \( G \) must force the second statement to be true as well. Therefore, we may suppose that \( p \) is an initial segment of \( G \) forcing the truth of both statements. Our forcing conditions will be the set \( \mathcal{P} = \{ p_i \in 2^{<\omega} \mid p \supseteq p_i \} \). We consider the set \( T = \{ r_j \mid r_j = \Psi^{p_j} \} \). Since \( p \) is enough to force totality, every element \( r_i \) in \( T \) will have a proper extension \( r_j \) in \( T \), so we can interpret \( T \) as an infinite recursively enumerable binary tree with \( X \) as one of its paths. If a real
Ψ^Z is not UD-random, the following statement \( P_{\alpha,\beta} \) must be true for some recursive sequence of distinct integers \( \langle a_i \rangle \) and the reals \( \alpha \) and \( \beta \) (where \( 0 \leq \alpha < \beta \leq 1 \)):

\[
P_{\alpha,\beta} : \ (\exists \varepsilon > 0)(\forall N)(\exists n > N) \left[ \frac{|\{i < n \mid \alpha < [a_i(\Psi^Z)] < \beta\}|}{n} - (\beta - \alpha) \geq \varepsilon \right].
\]

To show that \( X \) is not UD-random, it will be enough to set \( \alpha = \frac{1}{2} \), \( \beta = 1 \), and \( \varepsilon = \frac{1}{4} \) and build a recursive sequence \( \langle a_i \rangle \) of distinct integers such that for every \( r_i \in \mathbb{T} \), there is an extension \( r_j \) in \( T \) such that for some \( n > |r_i| \), we have:

\[
(\forall N)(\exists n > N) \left[ \frac{|\{i < n \mid \frac{1}{2} < [a_i(r_j)] < 1\}|}{n} < \frac{1}{4} \right]
\]

This guarantees that we can never force the fraction above to converge to the “correct” limit \( \frac{1}{2} \), and since \( G \) is 2-generic, \( G \) must force instead that \( \Psi^G \) is not UD-random.

To construct such a sequence \( \langle a_i \rangle \), we proceed in stages. At each stage, we have a finite sequence of \( a_i \) and add at least one element to it. At stage 0, our sequence is the empty sequence. At each subsequent stage \( s \) we deal with \( p_s \). Suppose we have already defined the finite initial segment \( \langle a_i \rangle_{i<M} \) of our \( a \)-sequence. Look for an \( r_j \) such that \( p_j \supseteq p_s \) and such that one of two conditions holds:

1. \( |\{i < M \mid \frac{1}{M} [a_i(r_j)] < 1\}| < \frac{1}{4} \) and \( r_j \) contains a 0 at some position \( k \) where \( a_i < 2^k \) for all \( i < M \).
2. There is an \( n \) such that \( |\{i < M \mid \frac{1}{M+n} [a_i(r_j)] < 1\}| < \frac{1}{4} \) and \( r_j \) contains a 0 at \( n \) positions \( k_0, k_2 \ldots k_{n-1} \) such that \( a_i < 2^{k_\ell} \) for all \( i < M, \ell \leq n \).

Such an \( r_j \) must exist. If Condition (1) holds, then we can find \( r_j \) as specified since we forced every extension of \( p \) to compute an \( r_j \in \mathbb{T} \) with infinitely many zeroes. Otherwise, we can find \( n \) as in Condition (2) and \( r_j \) with zeroes at \( n \) such positions for the same reason.

If Condition (1) holds, set \( a_M = 2^k \). This means that the fractional part of \( a_M(r_j) \) is less than \( \frac{1}{2} \) and we have \( |\{i < M+1 \mid \frac{1}{M+1} [a_i(r_j)] < 1\}| < \frac{1}{4} \) as well.

If Condition (2) holds, set \( a_{M+i} = 2^k \) for all \( 0 \leq i \leq n - 1 \). For each such \( i \), the fractional part of \( a_{M+i}(r_j) = 2^k(r_j) \) is less than \( \frac{1}{2} \) and we have \( |\{i < M+n \mid \frac{1}{M+n} [a_i(r_j)] < 1\}| < \frac{1}{4} \) as desired.

Now \( G \) must force either the \( \Pi^0_2 \) statement

\[
\Theta := (\forall N)(\exists n > N) \left[ \frac{|\{i < n \mid \frac{1}{2} < [a_i(\Psi^Z)] < 1\}|}{n} < \frac{1}{4} \right]
\]

or its negation. But \( G \) cannot force \( \neg \Theta \), so instead \( G \) must force \( \Theta \), so that \( \Psi^G = X \) cannot be UD-random. \( \square \)

This proof cannot be modified in an obvious way to show that no 1-generic \( tt \)-computes a UD-random. Were 2-genericity only necessary to force the totality of \( \Psi^X \), it would be possible, but the other statements in the problem are \( \Pi^0_2 \) as well.

4. Future Directions

It seems easier to prove negative results about UD-randoms than positive results. In part, this is due to the need to quantify over all intervals; however, this requirement has been minimized by Fact
1.2. Whether this fact has practical significance remains to be seen. However, the main difficulty in relating UD-randomness to other randomness notions is that UD-randomness is intrinsically a notion about tailsets: when we multiply a real by an integer, an initial segment of the real becomes irrelevant. In contrast, every other common randomness notion is defined in terms of initial segments: those that make up the components of a test or serve as inputs for the martingale or Kolmogorov complexity function.

**Question 4.1.** Is there a test, martingale, or string (e.g. Kolmogorov) complexity characterization of UD-randomness?

**Question 4.2.** Do the UD-random reals that are weakly 1-random coincide with the Schnorr random reals?

**Question 4.3.** Do the UD-random reals coincide with the Schnorr random reals in the hyperimmune-free degrees?

From the comment immediately preceding the questions, it is difficult to see how Question 4.1 could have a positive answer—at least, it is difficult to see how such a characterization could resemble that of other randomness notions in any meaningful way. The authors suspect that the answers to the second two questions are negative as well but have weaker grounds for saying so.

**References**


Department of Mathematics, Mailcode 4408, Southern Illinois University, 1245 Lincoln Drive, Carbondale, IL 62901, USA

E-mail address: wcalvert@siu.edu

Department of Mathematics, Room 306, Roosevelt Hall, Hofstra University, Hempstead, NY 11549-0114, USA

E-mail address: johanna.n.franklin@hofstra.edu